A Class of Estimators in Stratified Sampling with Two Auxiliary Variables

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SUMMARY

Following Srivastava [2], a general class of estimators for the finite population total utilizing the available knowledge on two auxiliary variables under a stratified sampling have been given.

Key words: Asymptotic variance, Auxiliary variable, Combined estimate, Separate estimate, Stratified sampling.

1. Introduction

Consider a finite population divided into L disjoint strata $S_1, S_2, ..., S_L$. Let Y_h and X_h be the totals of S_h in respect of the study variable y and an auxiliary variable x and the overall total $Y = \sum_h Y_h$ of y-values is to be estimated. Sampling within each stratum is done independently according to any probability sampling design. Given the sample s_h in S_h , let t_{hy} and t_{hx} be unbiased estimates of Y_h and X_h respectively, such that $V(t_{hy}) = \sigma_{hy}^2$, $V(t_{hx}) = \sigma_{hx}^2$ and Cov $(t_{hy}, t_{hx}) = \sigma_{hyx}$.

As is well-known, in a stratified sampling the auxiliary variable x can be incorporated in two different ways. If the overall total $X = \sum_h X_h$ of x-values is known, a combined estimate (e.g. ratio, product or regression estimate) for Y is made. In addition to the value of X, if the values of X_h (h = 1, 2, ..., L) are known, a separate estimator is built up from the stratum level estimators. But, in practice, the latter procedure yields better estimators than the former. In the present context, we use a second auxiliary variable to improve efficiency in the estimation of Y.

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2. Use of Second Auxiliary Variable

Sometimes even if X_h 's are known, information on a cheaply ascertainable variable z whose correlation with y may be less than that of x with y (i.e. $\rho_{yz} < \rho_{yx}$) is readily available. This type of situation may also be realizable with z is used as the stratification variable. For instance, in a crop survey if y, x, and z are respectively the yield of jute, area under jute and area under cultivation, information on the total cultivated area of each village can be obtained at a low cost. If the villages in a district are stratified on the basis of their total cultivated area, then information on z can be easily known from the district records.

Let t_{hz} be an unbiased estimate of Z_h (total of z-values in S_h) so that $t_z = \sum_h t_{hz}$ will be an unbiased estimate of Z (the overall total of z-values) with $V(t_z) = \sum_h \sigma_{hz}^2$ where $\sigma_{hz}^2 = V(t_{hz})$.

In this paper, following Srivastava [2], we develop a general class of estimators for Y using the covariates x and z simultaneously. We assume that, all the stratum level totals of x (i.e. X_h 's) are known but for the second auxiliary variable z only the overall total (i.e. Z) is known. It is also observed that many well-known, and some less-known but potentially interesting estimators belong to this class.

3. The Proposed Class of Estimators

For given $s_h \subset S_h$, (h = 1, 2, ..., L) let (t_{hy}, t_{hx}) assume values in a closed convex subspace, R_2 , of the two-dimensional real space containing the point (Y_h, X_h) . Then following Srivastava [2] a class of estimators for Y_h is defined by

$$\hat{Y}_h = g_h(t_{hy}, t_{hx})$$

where $g_h(t_{hy}, t_{hx})$ is a known function of t_{hy} and t_{hx} , independent of Y_h , such that $g_h(Y_h, X_h) = Y_h$ and satisfying the following regularity conditions:

- (i) The function $g_h(t_{hv}, t_{hr})$ is continuous in R₂
- (ii) The first and second order partial derivatives of $g_h(t_{hy}, t_{hx})$ exist and are also continuous in R_2 .

These conditions are assumed by Srivastava [2] for justifying Y_h to be a class of estimators for Y_h . On expanding $g_h(t_{hy}, t_{hx})$ about (Y_h, X_h) in a Taylor's series, we have to a first order of approximation

$$\hat{Y}_{h} \approx g_{h} (Y_{h}, X_{h}) + g_{h1} (t_{hy} - Y_{h}) + g_{h2} (t_{hx} - X_{h})$$
(3.1)

where g_{h1} and g_{h2} denote the first order partial derivatives of $g_h(t_{hy}, t_{hx})$ w.r.t. t_{hy} and t_{hx} respectively at (Y_h, X_h) . Noting that $g_h(Y_h, X_h) = Y_h$ and $g_{h1} = 1$, (3.1) can be rewritten as

$$\hat{Y}_h - Y_h = (t_{hy} - Y_h) + g_{h2}(t_{hx} - X)$$
 (3.2)

Thus, to a first order of approximation $E(\hat{Y}_b) = Y_b$ with asymptotic variance

$$V(Y_{h}) = \sigma_{hy}^{2} + 2g_{h2} \sigma_{hyx} + g_{h2}^{2} \sigma_{hx}^{2}$$
 (3.3)

Based on the above results, a class of separate estimators for Y may be defined by $t_s = \sum_h Y_h = \sum_h g_h(t_{hy}, t_{hx})$ with asymptotic variance

$$V(t_s) = \sum_{h} (\sigma_{hy}^2 + 2g_{h2} \sigma_{hyx} + g_{h2}^2 \sigma_{hx}^2)$$
 (3.4)

Whatever be the samples s_h (h = 1, 2, ..., L) and consequently an overall sample s (= $U s_h$) chosen, let (t_s , t_z) assume values in closed convex subspace, R_2' (say), of two-dimensional real space containing the point (Y, Z). Let $f(t_s, t_z)$ be a known function of t_s and t_z which may contain Z but independent of Y such that f(Y, Z) = Y, and also admitting the regularity conditions in R_2' .

It may be noted here that, R_2 and R_2' can be regarded as the yx and yz planes respectively of a three dimensional real subspace R_3 containing the points (Y_h, X_h, Z) and (Y, X, Z). Thus the points (Y_h, X_h) and (Y, Z) can also be represented by $(Y_h, X_h, 0)$ and (Y, 0, Z).

The proposed class of estimators of Y may be defined by

$$t_{G} = f(t_{s}, t_{z}) \tag{3.5}$$

Since there are only a finite number of possible samples, the expectation and variance of t_G exist under the condition (i). Expanding $f(t_*, t_*)$ about (Y, Z)

in a second order Taylor's series and taking expectation, we obtain the asymptotic variance of $t_{\rm G}$

$$V(t_G) = V(t_s) + 2f_2 Cov(t_s, t_z) + f_2^2 V(t_z)$$
(3.6)

where f_2 denotes the first order partial derivative of $f(t_s, t_z)$ w.r.t. t_z at (Y, Z).

Writing
$$t_{hz} = Z_h + (t_{hz} - Z_h)$$
 and using (3.2)
$$Cov(\hat{Y}_h, t_{hz}) \sim \sigma_{hyz} + g_{h2} \sigma_{hxz}$$
so that $Cov(t_s, t_z) = \sum_{h} Cov(\hat{Y}_h, t_{hz}) \sim \sum_{h} (\sigma_{hyz} + g_{h2} \sigma_{hxz})$

where
$$\sigma_{hvz} = Cov(t_{hv}, t_{hz})$$
 and $\sigma_{hxz} = Cov(t_{hx}, t_{hz})$

Finally, the formula for the asymptotic variance of t_G is obtained as

$$V(t_{G}) = \sum_{h} (\sigma_{hy}^{2} + 2g_{h2} \sigma_{hyx} + g_{h2}^{2} \sigma_{hx}^{2}) + f_{2}^{2} \sum_{h} \sigma_{hz}^{2} + 2f_{2} \sum_{h} (\sigma_{hyz} + g_{h2} \sigma_{hxz})$$
(3.7)

4. Some Observations and Remarks

4.1 If the second auxiliary variable z is not used, t_G reduces to t_S , i.e. a class generating a family of separate variety estimators. From (3.4) and (3.7) it follows that $V(t_G) \le V(t_S)$ if

$$f_2^2 \sum_{h} \sigma_{hz}^2 + 2f_2 \sum_{h} (\sigma_{hyz} + g_{h2} \sigma_{hxz}) \le 0$$

Thus, an estimator of t_G is more efficient than an estimator of t_S if

$$\beta_{yz} + \frac{\sum_{h} g_{hz} \sigma_{hxz}}{\sum_{h} \sigma_{hz}^2} \le -\frac{f_2}{2}$$
 (4.1)

where
$$\beta_{yz} = \sum_{h} \sigma_{hyz} / \sum_{h} \sigma_{hz}^{2}$$

This condition shows that there is a scope for improving upon the estimators based on one auxiliary variable x by using the second auxiliary variable z in stratified sampling.

4.2 If x is not used i.e. x-values are treated to be a non-zero constant, t_G reduces to a class of combined variety estimators represented by

$$t_c = f(t_y, t_z)$$
 with $t_y = \sum_h t_{hy}$

It may be noted here that, a different function satisfying the earlier regularity conditions can also be used to define t_c .

The asymptotic variance of t_c is

$$V(t_c) = \sum_{h} (\sigma_{hy}^2 + 2f_2\sigma_{hyz} + f_2^2 \sigma_{hz}^2)$$
 (4.2)

Thus, an estimator of t_G would be more efficient than that of t_c if

$$\beta_{\text{hyx}} + f_2 \beta_{\text{hzx}} \le -\frac{g_{\text{h2}}}{2}$$
 (h = 1, 2, ..., L) (4.3)

where $\beta_{hyx} = \sigma_{hyx}/\sigma_{hx}^2$, $\beta_{hzx} = \sigma_{hxz}/\sigma_{hx}^2$

4.3 The variance of t_G given in (3.7) is sought to be minimized subject to

$$g_{h2} = -(\beta_{hyx} + f_2\beta_{hzx}) = \hat{g}_{h2} \text{ (say)}$$

$$\frac{\sum_{h} \sigma_{hz}^2 (\beta_{hyz} - \beta_{hyx} \beta_{hxz})}{\sum_{h} \sigma_{hz}^2 (1 - \rho_{hxz}^2)} = \hat{f}_2 \text{ (say)}$$
(4.4)

and

where $\beta_{hyz} = \sigma_{hyz}/\sigma_{hz}^2$ and $\rho_{hxz} = \sigma_{hxz}/\sigma_{hx}\sigma_{hz}$

Thus, from (4.4), it is clear that, optimum values of g_{h2} and f_2 can not be determined uniquely. However, after obtaining an optimum value of f_2 , we can use this value to calculate the optimum values of g_{h2} . Utilizing these optimum values, the minimum asymptotic variance of f_G is given by

$$V_{min}(t_G) = \sum_{h} \sigma_{hy}^2 (1 - \rho_{hyx}^2 - B^2)$$
 (4.5)

where
$$B = \frac{\displaystyle\sum_{h} \sigma_{hy} \, \sigma_{hz} \, (\rho_{hyz} - \rho_{hyx} \, \rho_{hyz})}{\sqrt{\displaystyle\sum_{h} \sigma_{hy}^2} \, \sqrt{\displaystyle\sum_{h} \sigma_{hz}^2 \, (1 - \rho_{hxz}^2)}} \, \text{such that}$$

$$\rho_{hyx} = \, \sigma_{hyx} / \sigma_{hy} \, \sigma_{hx}, \, \rho_{hyz} = \, \sigma_{hyz} / \sigma_{hy} \, \sigma_{hz}$$

The estimator attaining this minimum variance is a regression-type estimator of the form

$$t_{RG} = \sum_{h} [t_{hy} - \hat{g}_{h2} (t_{hx} - X_{h})] - \hat{f}_{2} (t_{z} - Z)$$

studied earlier by Dalabehera and Sahoo [1]. This leads to an interesting result that, one cannot improve upon t_{RG} by using x and z for the situation under consideration

4.4 The suggested class provides us with an infinite number of estimators depending on proper choices of the functions g_h and f_s , and their asymptotic variances or mean square errors can be obtained from (3.7) by the substitution of corresponding values of g_{h2} and f_2 . For example, simple expansion estimator (without using x or z); separate variety ratio, product, difference and regression estimators using x; combined variety ratio, product, difference and regression estimators using z are particular members of the class. It is also interesting to note that the classes of estimators represented by t_s and t_c may be identified as subclasses of the class of estimators represented by t_s .

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